



# Merging uncertainty sets via majority vote

Matteo Gasparin<sup>1</sup> Aaditya Ramdas<sup>2</sup>

<sup>1</sup>University of Padova <sup>2</sup>Carnegie Mellon University



## Introduction

- In statistics, uncertainty is commonly captured through uncertainty sets (i.e., confidence intervals or prediction sets).
- In certain scenarios, different (dependent) uncertainty sets are generated by different agents.
- Some examples are conformal prediction intervals based on different algorithms or confidence intervals for a parameter of interest based on different methods.
- How should we combine  $K$  **arbitrarily dependent uncertainty sets**?

## Problem statement

- Input:**  $\mathcal{C}_1, \dots, \mathcal{C}_K$  are  $K \geq 2$  arbitrarily dependent uncertainty sets satisfying  $\mathbb{P}(c \in \mathcal{C}_k) \geq 1 - \alpha$ , for all  $k = 1, \dots, K$ .
- Output:** a single set that combines them in a black-box manner.

Two important quantities to consider: **coverage** and **size**.

Two naive solutions:

- $\bigcup_{k=1}^K \mathcal{C}_k$  has coverage  $1 - \alpha$ , but it is too conservative.
- $\bigcap_{k=1}^K \mathcal{C}_k$  has coverage  $1 - K\alpha$ , but it is too anti-conservative.

## Majority vote

Include all the points that are contained in at least half of the sets.

$$\mathcal{C}^M := \left\{ s \in \mathcal{S} : \frac{1}{K} \sum_{k=1}^K 1\{s \in \mathcal{C}_k\} > \frac{1}{2} \right\}.$$

Using Markov's inequality:  $\mathbb{P}(c \in \mathcal{C}^M) \geq 1 - 2\alpha$ .

In addition,

$$m(\mathcal{C}^M) \leq \frac{2}{K} \sum_{k=1}^K m(\mathcal{C}_k),$$

where  $m(\cdot)$  denotes the Lebesgue measure of a set.

## Summary of the main results

- Majority vote is a good way to merge uncertainty sets.
- Improvements achieved through **randomization** and **exchangeability**.
- Drawback: In some cases (rarely in sims), the output is a union of intervals.
- The method can be used to derandomize statistical procedures based on **data splitting**.

## Adding prior information

If there is a belief that certain agents are more accurate  $\rightarrow$  incorporate prior information through a prior distribution  $w = (w_1, \dots, w_k)$  over the agents.

Weighted majority vote:

$$\mathcal{C}^W := \left\{ s \in \mathcal{S} : \sum_{k=1}^K w_k 1\{s \in \mathcal{C}_k\} > \frac{1}{2} \right\}.$$

In this case:  $\mathbb{P}(c \in \mathcal{C}^W) \geq 1 - 2\alpha$  and  $m(\mathcal{C}^W) \leq 2 \sum_{k=1}^K w_k m(\mathcal{C}_k)$ .

## Improving majority vote with randomization

Let  $u \sim \text{Unif}(0, 1)$ , independent of all the data. Define

$$\mathcal{C}^R := \left\{ s \in \mathcal{S} : \sum_{k=1}^K w_k 1\{s \in \mathcal{C}_k\} > \frac{1}{2} + u/2 \right\}.$$

We obtain that  $\mathcal{C}^R \subseteq \mathcal{C}^W$  and  $\mathbb{P}(c \in \mathcal{C}^R) \geq 1 - 2\alpha$ .

The proof is based on the uniformly-randomized Markov inequality.

Another possibility is to define the set  $\mathcal{C}^U$  with a completely random threshold  $u$ , in this case  $\mathbb{P}(c \in \mathcal{C}^U) \geq 1 - \alpha$ .

## Merging exchangeable sets

- When  $\mathcal{C}_1, \dots, \mathcal{C}_K$  are exchangeable, it is possible to obtain something better than a naive majority vote.
- We denote  $\mathcal{C}^M(1 : K) = \mathcal{C}^M$  to highlight that it is based on the majority vote of sets  $\mathcal{C}_1, \dots, \mathcal{C}_K$ .

We define

$$\mathcal{C}^E := \bigcap_{k=1}^K \mathcal{C}^M(1 : k).$$

By definition  $\mathcal{C}^E \subseteq \mathcal{C}^M$ , in addition  $\mathbb{P}(c \in \mathcal{C}^E) \geq 1 - 2\alpha$ .

A simple way to improve the majority vote for arbitrarily dependent sets: process them in a random order ( $\mathcal{C}^\pi$ ).

## Derandomizing statistical procedures

It can be used also for **point estimators**.

**Theorem:** Suppose  $\hat{\theta}_1, \dots, \hat{\theta}_K$  are  $K$  univariate point estimators of  $\theta$  that are based using  $n$  data points and satisfy a high probability concentration bound

$$\mathbb{P}(|\hat{\theta}_k - \theta| \leq w(n, \alpha)) \geq 1 - \alpha,$$

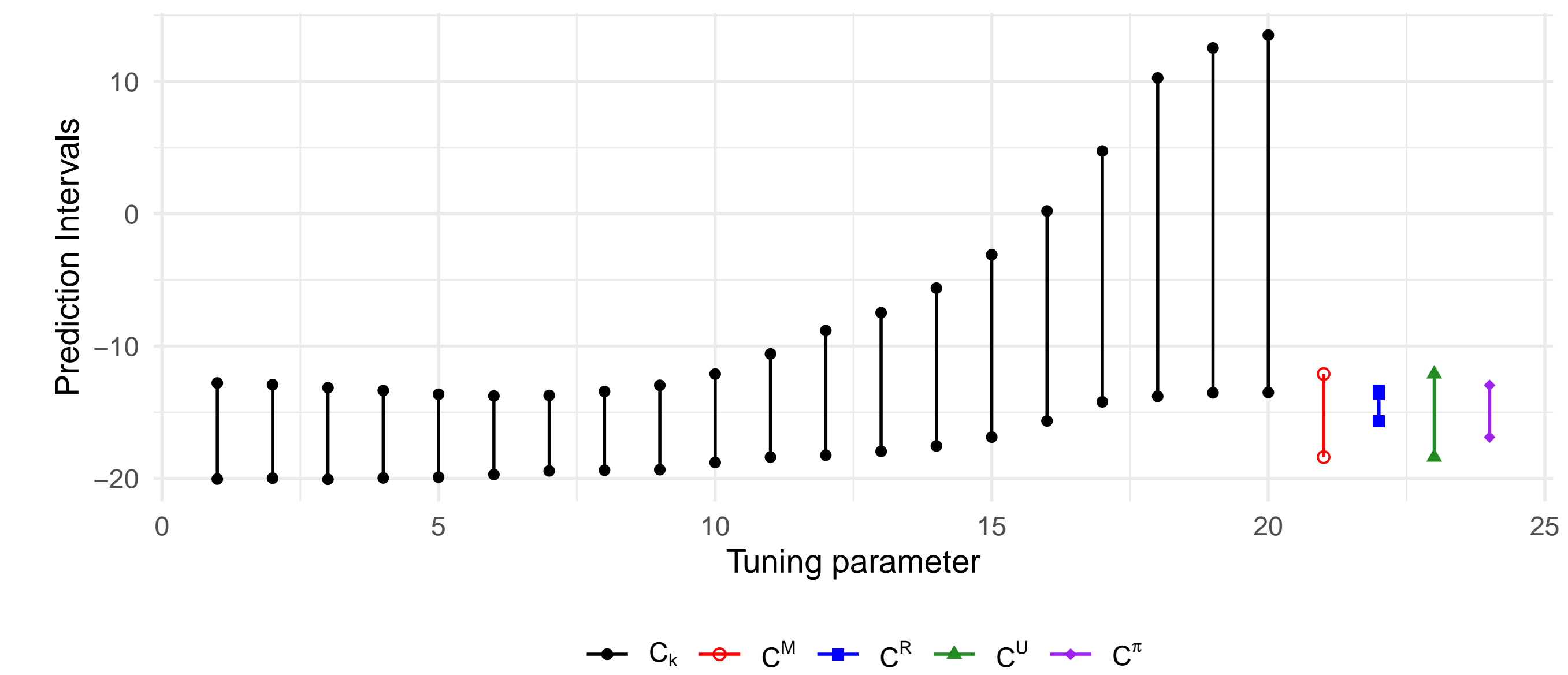
for some function  $w$ . Then, their median  $\theta_{(\lceil K/2 \rceil)}$  satisfies

$$\mathbb{P}(|\hat{\theta}_{(\lceil K/2 \rceil)} - \theta| \leq w(n, \alpha)) \geq 1 - 2\alpha. \quad (1)$$

Further, if  $\hat{\theta}_1, \dots, \hat{\theta}_K, \dots$  are exchangeable, then (1) is uniformly valid.

## Example: conformal prediction with lasso

Fit lasso regression to data, with different penalty parameters  $\lambda$  and  $\alpha = 0.05$ .

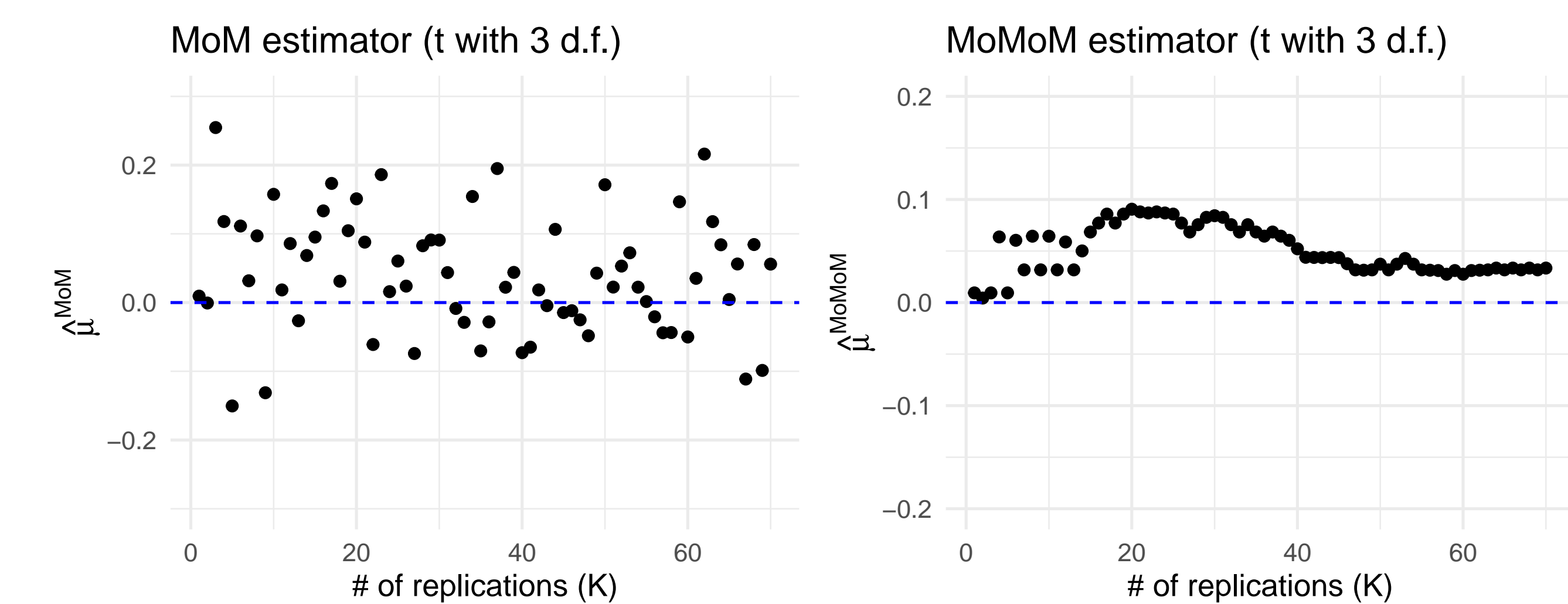


Randomized sets used  $u = 1/2$  for visualization.  
Coverage:  $\mathcal{C}^M = 0.97$ ,  $\mathcal{C}^R = 0.92$ ,  $\mathcal{C}^U = 0.96$ ,  $\mathcal{C}^\pi = 0.93$ .

## Derandomizing MoM (Median-of-Means)

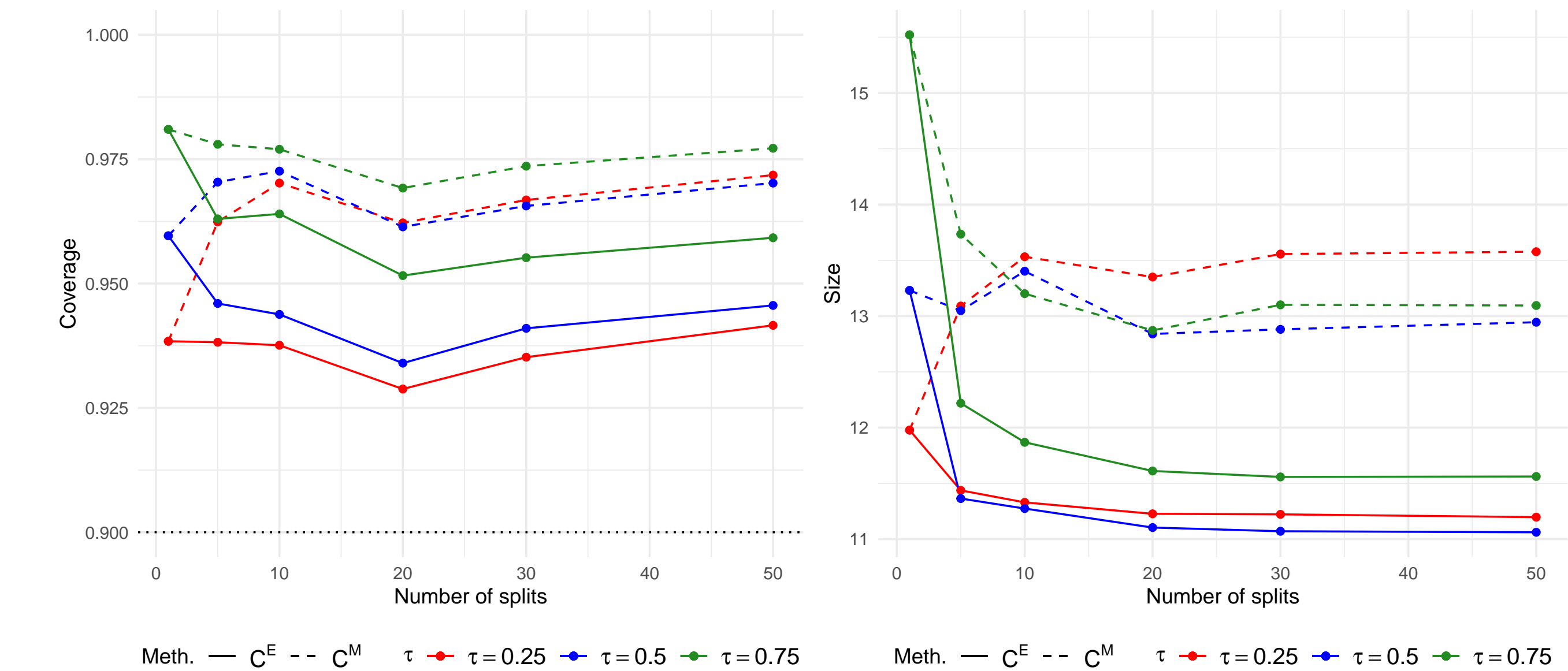
$\hat{\mu}^{\text{MoM}}$ : Estimator of the mean for  $X_1, \dots, X_n \stackrel{iid}{\sim} P$  based on data-splitting.

$$\hat{\mu}^{\text{MoMoM}} := \text{median}(\hat{\mu}_1^{\text{MoM}}, \dots, \hat{\mu}_K^{\text{MoM}})$$



## Multi-split conformal inference

Construct  $K$  split conformal prediction intervals + (exchangeable) majority vote.



$\mathcal{C}^E$ : smaller sets and coverage closer to the level  $1 - \alpha = 0.9$ .